Effective monitoring of carbon emissions from industrial sector using statistical process control

Shamsuzzaman, Mohammad; Shamsuzzoha, Ahm; Maged, Ahmed; Haridy, Salah; Bashir, Hamdi; Karim, Azharul

Published in:
Applied Energy

Published: 15/10/2021

Document Version:
Final Published version, also known as Publisher’s PDF, Publisher’s Final version or Version of Record

License:
CC BY

Publication record in CityU Scholars:
Go to record

Published version (DOI):
10.1016/j.apenergy.2021.117352

Publication details:

Citing this paper
Please note that where the full-text provided on CityU Scholars is the Post-print version (also known as Accepted Author Manuscript, Peer-reviewed or Author Final version), it may differ from the Final Published version. When citing, ensure that you check and use the publisher's definitive version for pagination and other details.

General rights
Copyright for the publications made accessible via the CityU Scholars portal is retained by the author(s) and/or other copyright owners and it is a condition of accessing these publications that users recognise and abide by the legal requirements associated with these rights. Users may not further distribute the material or use it for any profit-making activity or commercial gain.

Publisher permission
Permission for previously published items are in accordance with publisher's copyright policies sourced from the SHERPA RoMEO database. Links to full text versions (either Published or Post-print) are only available if corresponding publishers allow open access.

Take down policy
Contact lbscholars@cityu.edu.hk if you believe that this document breaches copyright and provide us with details. We will remove access to the work immediately and investigate your claim.
Effective monitoring of carbon emissions from industrial sector using statistical process control

Mohammad Shamsuzzaman a,*, Ahm Shamsuzzoha b, Ahmed Maged c, d, Salah Haridy a, d, Hamdi Bashir a, d, Azharul Karim e

a Department of Industrial Engineering and Engineering Management, Sustainable Engineering Asset Management (SEAM) Research Group, College of Engineering, University of Sharjah, United Arab Emirates
b School of Technology and Innovations and Digital Economy Research Platform, University of Vaasa, Vaasa 65101, Finland
c Department of Systems Engineering and Engineering Management, City University of Hong Kong, Tat Chee Avenue, Kowloon, Hong Kong
d Department of Mechanical Engineering, Benha Faculty of Engineering, Benha University, Egypt
e Department of Mechanical Engineering, Leader, Energy & Drying Research Group, Science and Engineering Faculty, Queensland University of Technology, 2 George St, Brisbane, QLD 4000, Australia

HIGHLIGHTS

• A scheme for effective monitoring and controlling of carbon emissions is proposed.
• The scheme is optimized for detecting increasing shifts in carbon emissions.
• Effectiveness of the proposed scheme is investigated under different scenarios.
• Continuous monitoring of carbon emission reduces the related costs significantly.
• Valuable insights are provided for designing the proposed monitoring scheme.

ARTICLE INFO

Keywords:
Energy consumption Carbon emissions Industry Environmental quality management Statistical process monitoring Economic Shewhart-EWMA scheme Optimization design

ABSTRACT

The industrial sector is considered one of the fastest-growing sources of greenhouse gases, due to the excessive consumption of energy required to cope with the growing production of energy exhaustive products. The statistical process monitoring (SPM) can be an effective tool for monitoring and controlling carbon emissions from industries. This article presents an economic-statistical design of the combined Shewhart $\bar{X}$ and exponentially weighted moving average (EWMA) scheme ($\bar{X}$&EWMA scheme) for monitoring carbon emissions from industries to allow prompt action for controlling excessive emissions. The parameters of the proposed SPM scheme have been optimized for minimizing the expected total cost, including cost from carbon emissions and operational costs of the SPM scheme. The design of the $\bar{X}$&EWMA scheme has been optimized considering a wide range of shifts in the mean of the emission process, and ensuring that the constraints on inspection rate, sample size, and false alarm rate are all satisfied. Comparative studies showed that the optimal $\bar{X}$&EWMA scheme reduced the expected total cost by about 40%, 77%, and 28% compared with the basic $\bar{X}$, EWMA, and $\bar{X}$&EWMA schemes, respectively. The impact of the design parameters on the effectiveness of the proposed SPM scheme has also been investigated by sensitivity analysis. Finally, the application of the proposed SPM scheme is demonstrated by using real data for carbon emissions from different industrial facilities. This study is expected to considerably reduce the cost owing to excessive carbon emissions from industries and widen the literature on the utilization of SPM tools in managing the quality of the environment.

* Corresponding author.
E-mail addresses: mshamsuzzaman@sharjah.ac.ae (M. Shamsuzzaman), ahm.shamsuzzoha@uwasa.fi (A. Shamsuzzoha), amaged2-c@my.cityu.edu.hk (A. Maged), sharidy@sharjah.ac.ae (S. Haridy), hbashir@sharjah.ac.ae (H. Bashir), azharul.karim@qut.edu.au (A. Karim).

https://doi.org/10.1016/j.apenergy.2021.117352
Received 15 March 2021; Received in revised form 29 May 2021; Accepted 9 June 2021
Available online 12 July 2021
0306-2619/© 2021 The Authors. Published by Elsevier Ltd. This is an open access article under the CC BY license (http://creativecommons.org/licenses/by/4.0/).
1. Introduction

Environmental degradation is considered as one of the most critical issues by today’s researchers, professionals, and policymakers. Human activities increase the emission of heat-trapping gases, known as greenhouse gases (GHGs) that cause global warming and ecological imbalances [1,2]. Uprety et al. [3] reported that the concentration of GHGs in the Earth’s atmosphere has significantly increased since the pre-industrial era of 1850. This increase of GHGs emissions is changing the Earth’s climate, leading to various catastrophic events such as floods, earthquakes, droughts, and the deterioration of the polar regions [1]. According to the Kyoto Protocol, carbon dioxide (CO₂) is one of the six major GHGs that potentially influence the climate [4].

Since 1990, the industrial sector has grown by 174% and is considered the fastest-growing sources of GHGs [5]. Carbon emissions from different industrial sectors are characterized by a wide range of emission quantities, depending on the type of industrial sector, the type of technology used, and the energy source. Although energy consumed by different industrial sectors has decreased in recent years, the total energy use has still increased due to production growth and the increase of energy exhaustible industrial products [6]. The International Energy Agency (IEA) assumes that industrial energy use will continue to increase until it approximately doubles in 2050, as compared to the consumption levels in 2009. As a result, the industrial CO₂ emissions are expected to increase by 45–65% [6]. Chontanawat [7] investigated the dynamic relationship between energy consumption and carbon emissions using co-integration and causality models, and concluded that energy consumption causes carbon emissions, implying their increases are directly proportional to each other. Thus, monitoring carbon emissions is the key towards encouraging households, businesses, and industries to use energy-efficient products as well as clean energy.

Many industrialized countries have imposed environmental legislations or carbon taxes and/or implemented the cap-and-trade system to control fossil fuel emissions and promote cleaner energy [8–10]. The carbon tax is a surcharge applied on GHGs emissions, mainly from burning fossil fuels. For instance, Sweden has imposed a carbon tax since 1991 to minimize GHGs emissions, and the federal government of Canada has been enforcing rules and regulations nationwide to ensure all provinces have a carbon fee in place. On the contrary, in a cap-and-trade system, governments put a threshold or cap on the average amount of carbon emissions from an industry. The United States and the European Union have been successfully implementing the cap-and-trade system to meet the commitments toward controlling GHGs emissions [11,12].

However, researchers are continuously exploring varying techniques to find the most effective way to control carbon emissions. Chen et al. [13] proposed an inexact multi-criteria decision-making model for ensuring the optimal lifecycle economics and GHGs emissions under uncertainty. To promote sustainable development of human society, the transition of the global energy system from high-carbon to low-carbon energy resources, such as shale gas, is essential. Chen et al. [14] developed a multi-level programming model for lifecycle assessment of GHGs emissions and water-energy optimization for a shale gas supply chain. Similarly, He et al. [15] evaluated shale gas resources and their corresponding environmental implications under uncertainty. Abeydeera et al. [16] emphasized on monitoring and documenting the amount of carbon emissions at various levels (product, organization, city, and national) and the objective of formulating the necessary strategies to manage the quality of the environment. Likewise, [17] and [18] developed systems for monitoring and assessing the environmental performance of the real estate sector in Sweden via environmental indicators. They concluded that the energy and emissions of buildings can be estimated using time series models. According to [19], analyzing the patterns of the recently monitored data of carbon emissions can be very beneficial to efficiently assess current and future carbon emission trends. Thereafter, many studies have been directed to evaluate the emission rates from different industries, such as manufacturing [20] and energy [21]. In addition, statistical process monitoring (SPM) schemes can be used successfully to continuously monitor the emissions data and identify unusual changes in a timely fashion [22–24].

The continuous monitoring of carbon emissions from industries using SPM schemes can provide several benefits. At the industrial level, it can help in identifying excessive emissions at an early stage, and thus ensure that appropriate action can be taken in advance to control them, which in turn can minimize the expected total cost including emission-related and operational costs of the SPM scheme. For policymakers, it can assist in (i) evaluating whether the emissions are within the regulatory limit (e.g., carbon-cap as specified by the government) or at a high risk of non-compliance, (ii) adjusting the control parameters in a systematic way to avoid non-compliance, (iii) monitoring and measuring the impact and related costs of emissions on the environment, (iv) establishing guidelines for evaluating real-time emissions against the targeted emissions and regulatory requirements, and (v) deciding which facility needs more frequent inspection, based on the frequency of the signal produced by the SPM schemes. Most importantly, SPM schemes can help decision-makers set an appropriate amount of emission fee (i.e., carbon tax).

The remainder of the paper is organized as follows. Section 2 reviews the relevant articles, identifies the research gaps, and highlights the contribution of the paper. Section 3 develops the model for the optimization design of the proposed SPM scheme. Section 4 discusses the results of numerical studies conducted to evaluate the performance of the proposed SPM scheme under different operational scenarios. Section 5 illustrates the design and application of the proposed SPM scheme through a case study. Finally, the conclusions and future research directions are discussed in Section 6.

2. Literature review

The SPM tools have been mainly applied for measuring and controlling the quality of products in manufacturing industries for over 50 years, where the SPM chart is commonly used for monitoring a manufacturing process behavior over time to identify any unusual changes or trends, which ultimately helps in reducing the waste and improving the quality of the product [25]. The widespread application of SPM charts in manufacturing is mainly due to the fact that the quality characteristics (e.g., dimensions of a product) in a manufacturing setting can easily be defined and measured. Moreover, the flow of the products throughout most of the manufacturing processes can easily be tracked and controlled. On the other hand, the application of SPM tools in non-manufacturing sectors is really challenging because of the invisible work processes, lack of data, and difficulties in standardizing and measuring the quality characteristics. Although the application of SPM charts is comparatively less in non-manufacturing sectors, its adoption is growing rapidly because of the significant improvement in data acquisition and powerful computing systems in the recent years. One such sector is the environmental quality management (EQM). The quality of environment (e.g., quality of ambient air) can be affected by different sources, including carbon emissions from industrial facilities. Few researchers have proposed the application of SPM charts for controlling and managing the quality of environmental processes through the effective monitoring of environmental characteristics, such as pollutants discharged from different industries into the environment (for instance, see [26,27]). Madu [26] explained how SPM schemes can be used for environmental monitoring, while [22] designed a traditional cumulative sum (CUSUM) chart for monitoring the nitrate concentration blank measurement data. Furthermore, they used the process capability indices to evaluate the environmental performance of the nitrate blank process to avoid associated risks. Pan and Chen [23] designed an economic CUSUM scheme based on Duncan’s model and compared its performance with that of the X scheme for monitoring liquid (zinc)
waste and industrial pollutants discharged into a river. Leiva et al. [28] designed an attribute control chart for monitoring environmental risks due to dangerous pollutants present in the air, and the performance of the proposed methodology was investigated via simulation study. Similar to [28], Marchant et al. [29] proposed a methodology for monitoring particulate matter pollutants present in the environment using bivariate SPM charts. Capezza et al. [30] discussed traditional multivariate techniques for monitoring the total CO₂ emissions from a cruise ship, on different voyages, to detect anomalous occurrences.

The abovementioned SPM schemes are designed for monitoring either a single value or a few specific values of process shift. However, in almost all real applications, predicting the process shift is extremely cumbersome because the size δ of a shift in the process mean is a random variable that varies from time to time [31]. Consequently, an SPM scheme established considering a single value or a few specific values of δ may not satisfactorily capture the real characteristics of the process effectively. On the contrary, if data on δ are collected, the distribution of δ can be estimated and an optimal SPM scheme can be designed so that its effectiveness can be enhanced over a wide range of δ rather than a specific shift point. In addition, most of the abovementioned models assume that the quality characteristics X to be monitored are normally distributed, which is not always the case in environmental pollution processes. In most of the real applications, the environmental data is non-normal, and thus the traditional SPM schemes cannot be used directly for monitoring them. Liu and Xue [24] proposed a cost-based exponentially weighted moving average (EWMA) scheme (known as ML-EWMA chart) for monitoring the non-normal environmental data, assuming a random shift δ in the environmental pollution process. The proposed model minimizes only the quality loss experienced by an environmental pollution process based on Taguchi’s loss function. However, the primary goal of implementing an economic SPM scheme is to minimize the expected total cost, including the cost due to quality loss and operational costs of the SPM scheme. Several extensions to the pioneering economic design of X scheme, developed by [32], have been proposed (for instance, see [33–37]). Although the economic design of an SPM scheme is popular, it suffers from poor statistical properties (e.g., high false alarm rate). Therefore, several scholars have developed economic-statistical designs of an SPM scheme to reduce the false alarm rate (for instance, see [36,38–40]).

It is well-known that the traditional X scheme is a better choice for detecting large process shifts, whereas the EWMA and CUSUM schemes are mainly used for detecting small process shifts [25]. The effectiveness of the EWMA scheme is comparable to that of the CUSUM scheme; however, the former is easier to design and operate [25]. An SPM scheme combining both X and EWMA charts can enhance the performance of the monitoring scheme for detecting both small and large shifts in the environmental pollution processes. This study presents an optimization model for the economic-statistical design of the combined X& EWMA scheme for monitoring carbon emissions from industrial sectors, considering random shifts in the emission process. The contribution of the proposed study is summarized as follows: (i) The proposed model optimizes the charting parameters including the sample size, sampling interval, weighting parameter, and control limits of the combined X& EWMA scheme, and, in the meantime, ensures that no extra resources for operating the SPs scheme will be necessary. (ii) The proposed SPM scheme minimizes the expected total cost including the cost due to quality loss in the emission process and the operational cost of the SPM scheme. (iii) The performance of the proposed X&EWMA scheme is compared to basic X, EWMA and X&EWMA schemes. The study shows that the proposed combined X&EWMA scheme is significantly superior to its competitors for monitoring an environmental process. (iv) The performance of the proposed X&EWMA scheme is investigated extensively under different operational scenarios to help practitioners identify the optimal charting parameters using a computer program, available upon request. (v) The design and application of the proposed SPM scheme are illustrated by a real case study to promote its practical use.

3. Model development

3.1. Assumptions

Formulating the model proposed in this article involves the following assumptions:

1. The emission process begins from an in-control (IC) condition. The carbon emission variable x is independent and has normal distribution, with IC mean μ₀ and standard deviation (SD) σ₀. An assignable cause will alter the IC mean μ₀ to out-of-control (OOC) mean μ₁:

\[ μ₁ = μ₀ + δσ₀ \]  

where δ is the size of the shift in the mean value of the carbon emission process, experienced by an assignable cause, if the emission process is in the IC state, δ = 0. To simplify the process of designing the model, the shift in the mean of the emission process is not considered in this study (i.e., δ = 0).

2. The shift of size δ in the mean value of the carbon emission process is characterized by a Rayleigh distribution. This distribution is well-accepted in the SPM research community as a realistic representative of the distribution of the process shift [41–43].

3. The OOC state occurs owing to a single assignable cause in the emission process. The incidence of the assignable cause is assumed to follow a homogenous Poisson process, with mean λ₀ (i.e., the length of the IC state of the emission process follows an exponential distribution with a mean of 1/λ₀). This is a critical assumption, however, such assumption substantially simplifies the process of designing the economic model [25].

4. The carbon emission process continues during identifying and fixing the assignable cause.

3.2. Notations

The notations used in this study and their definitions are presented in Table 1.

<table>
<thead>
<tr>
<th>Symbol</th>
<th>Definition</th>
</tr>
</thead>
<tbody>
<tr>
<td>λ</td>
<td>The EWMA weighting factor.</td>
</tr>
<tr>
<td>n</td>
<td>Sample size</td>
</tr>
<tr>
<td>h</td>
<td>Sampling interval</td>
</tr>
<tr>
<td>UCL</td>
<td>Upper control limit of the X chart.</td>
</tr>
<tr>
<td>H</td>
<td>Upper control limit of the EWMA chart.</td>
</tr>
<tr>
<td>μ₀</td>
<td>Mean amount of carbon emissions during the IC state of the emission process.</td>
</tr>
<tr>
<td>σ₀</td>
<td>Standard deviation of the amount of carbon emissions during the IC state of the emission process.</td>
</tr>
<tr>
<td>USL</td>
<td>Upper specification limit of the amount of carbon emissions.</td>
</tr>
<tr>
<td>Q</td>
<td>Amount of carbon emissions per unit time.</td>
</tr>
<tr>
<td>λ₀</td>
<td>Incidence rate of the assignable cause.</td>
</tr>
<tr>
<td>O</td>
<td>Maximum number of carbon emission data inspected per unit time (i.e., maximum permissible inspection rate).</td>
</tr>
<tr>
<td>μ₁</td>
<td>Mean of the δ values in the carbon emission process.</td>
</tr>
<tr>
<td>g</td>
<td>Time required to estimate and test an observed data of a sample of carbon emission.</td>
</tr>
<tr>
<td>tₐ</td>
<td>Time length from an OOC state to the identification and fixation of the assignable cause.</td>
</tr>
<tr>
<td>δ₀</td>
<td>Minimum allowable IC ATS₀.</td>
</tr>
<tr>
<td>a₁</td>
<td>Fixed part of the sampling cost.</td>
</tr>
<tr>
<td>a₂</td>
<td>Variable part of the sampling cost.</td>
</tr>
<tr>
<td>a₃</td>
<td>Cost of detecting and dissecting an assignable cause.</td>
</tr>
<tr>
<td>a₄</td>
<td>Cost of investigating a false alarm.</td>
</tr>
<tr>
<td>Cₓ</td>
<td>The average penalty cost for an out-of-specification amount of carbon emissions.</td>
</tr>
</tbody>
</table>
amount of carbon emissions may be the deterioration of the equipment proposed in this research. The process parameters (\( \mu_0, \sigma_0 \)) can be estimated from the data observed in the pilot runs or process capability studies. The value of the USL can be decided based on the permissible amount of carbon emissions or carbon-cap; the amount of carbon emissions is supposed to not exceed the USL. The value of \( Q \) may be estimated from the company’s historical records of energy consumption per unit time. The value of the rate of occurrence of assignable cause \( \lambda_a \) can be estimated based on the historical records of OOC cases. The presence of the assignable cause incurs an excessive amount of carbon emissions is supposed to not exceed the USL, the proposed X&EWMA scheme is optimized for identifying the increasing shifts in the emission process. Consequently, an upper-sided EWMA scheme and an upper-sided X scheme have been combined. The aforementioned model optimizes \( \lambda, n, h, UCL, H \), only \( n \) and \( \lambda \) are independent. The value of \( h \) is determined such that the constraint on \( o \) is satisfied:

\[
h = n/O.
\]

UCL and \( H \) are determined such that the constraint on IC ATS\(_0\) (constraint (4)) is satisfied. The objective function ETC is calculated as follows:

\[
ETC = \int_0^\infty [TC(\delta) \cdot f(\delta)]d\delta
\]

where TC(\( \delta \)) is the total cost incurred owing to carbon emissions per unit time of an operational cycle for a given shift of size \( \delta \) in the carbon emission process. The calculation of TC(\( \delta \)) is explained in the following sections. The probability density function \( f(\delta) \) in Eq. (7) is obtained from Rayleigh distribution, as expressed below:

\[
f(\delta) = \frac{\delta}{\delta^2} \exp\left(-\frac{\delta^2}{2\delta^2}\right)
\]

It can be seen that the probability density function \( f(\delta) \) of the Rayleigh distribution is modeled by a single variable—the mean value \( \mu_0 \) of \( \delta \). It can be noted that the data on \( \delta \) can be obtained through a three-phase statistical process control (SPC) scenario, as suggested by [42].

3.3.2. Estimation of the total cost, TC(\( \delta \))

For any given shift of size \( \delta \) in the emission process, the total cost per unit time of an operational cycle, TC(\( \delta \)), is calculated from the ratio of expected cost, EC(\( \delta \)), to the expected length, EL(\( \delta \)), of the operational cycle.

3.3.2.1. Estimation of the expected length EL(\( \delta \)) of an operational cycle.

The time length \( L \) of an operational cycle is the time period from the beginning (or restoration) of the emission process to the identification and fixation of an assignable cause. This \( L \) comprises four time components: the IC time period (\( t_1 \)); the OOC time period (\( t_2 \)), the amount of time spent in taking a sample (size \( n \)) of carbon emission data and analyzing it (\( t_3 \)); and the time length from an OOC state to the identification and fixation of an assignable cause (\( t_4 \)). These four time components are random variables and only their expected values can be obtained.

As indicated earlier, the time between incidences of the assignable causes is assumed to be an exponential distribution with incidence rate \( \lambda_a \); therefore, the mean time between incidences of the assignable causes (i.e., mean time length of an IC state) is as follows:

\[
t_1 = 1/\lambda_a.
\]

If an assignable cause occurs between two consecutive samples, then the time component \( t_2 \) can be estimated as follows [25,32]:

\[
\alpha \leq O, n \leq n_{\text{max}}.
\]
\[ t_1 = ATS_1(\delta) - \tau = ATS_3(\delta) = \left( \frac{h}{3} - \frac{\lambda \delta^2}{12} \right) \]  

(10)

where \( \tau \) is the expected time of incidence of the process shift (of size \( \delta \)) between the \( j \)th and \((j + 1)\)th sample, given that the shift occurs during this interval. The expected value of the time period \( t_3 \) can be estimated in a straightforward manner, based on \( n \) and \( g \).

\[ t_3 = \frac{g \cdot n}{\lambda} \]  

(11)

Finally, the time period from an OOC state (owing to a process shift of size \( \delta \)) to the identification and fixation of an assignable cause \( t_4 \) can be approximated, based on the historical records of OOC cases.

The expected time length, \( EL(\delta) \), can now be calculated based on the time components \( t_1, t_2, t_3, \) and \( t_4 \).

\[ EL(\delta) = t_1 + t_2 + t_3 + t_4 = \frac{1}{\lambda_0} + ATS_3(\delta) - \left( \frac{h}{3} - \frac{\lambda \delta^2}{12} \right) + gn + t_4 \]  

(12)

3.3.2.2. Estimation of the expected cost \( EC(\delta) \) of an operational cycle. The primary goal of employing SPM tools is to optimize (i.e., minimize) the ETC that includes the quality cost (i.e., cost incurred owing to carbon emissions) and the operational cost of the SPM scheme. The quality cost in an operational cycle \( (C_2) \) can be estimated by utilizing the Taguchi’s loss function concept [44]. The operational cost of the SPM scheme, in this case, as the cost of sampling and estimating carbon emission data \( (C_2) \), cost of examining a false alarm \( (C_3) \), and cost of detecting and dissecting an assignable cause \( (C_4) \) in an operational cycle, can be estimated on the basis of the cost parameters, specified in the model developed by [32]. The quality cost \( C_1 \), defined as the cost incurred owing to a shift of size \( \delta \) in the carbon emission process, can be estimated from the basis of Taguchi’s loss function concept [44].

\[ C_1 = \left[ EL(\delta) - \frac{1}{\lambda_0} \right] Q \cdot K \cdot (\sigma_0^2 + \delta^2 \sigma_0^2) \]  

(13)

\[ K = \frac{C_1}{(USL - \mu_0)^2} \]

Here, \( 1/\lambda_0 \) is the time length of the IC period, \[ EL(\delta) - 1/\lambda_0 \] is the time length of the OOC period owing to a shift of size \( \delta \), and \( K \) is the cost coefficient, estimated based on the cost component \( C_1 \) associated with the USL (carbon-cap).

The expected cost of sampling and estimating the carbon emission data \( C_2 \), can be estimated based on the fixed \( (a_1) \) and variable \( (a_2) \) sampling cost components.

\[ C_2 = \frac{(a_1 + a_2 n) \cdot EL(\delta)}{h} \]  

(14)

The expected cost of investigating a false alarm in an operational cycle \( (C_3) \), can be determined based on the time length of the IC period \( 1/\lambda_0 \), the IC \( ATS_0 \), and the cost of examining a false alarm, \( a_4 \).

\[ C_3 = \frac{a_4}{\lambda_0 \cdot ATS_0} \]  

(15)

Thus, the expected cost incurred owing to a shift of size \( \delta \) in the emission process in an operation cycle, \( EC(\delta) \), can be obtained by adding all the cost components, \( C_1, C_2, C_3, \) and \( a_3 \).

\[ EC(\delta) = C_1 + C_2 + C_3 + a_3 \]

\[ = \left[ EL(\delta) - \frac{1}{\lambda_0} \right] Q \cdot K \cdot (\sigma_0^2 + \delta^2 \sigma_0^2) + \frac{(a_1 + a_2 n) \cdot EL(\delta)}{h} + \frac{a_4}{\lambda_0 \cdot ATS_0} + a_3 \]  

(16)

Finally, the total cost incurred owing to carbon emissions, per unit time of an operational cycle, for any given value of \( \delta \), \( TC(\delta) \), can be found as follows:

\[ TC(\delta) = \frac{EC(\delta)}{EL(\delta)} \]  

(17)

In summary, for any given set of values of the process parameters \( (\lambda, O, \mu_0, g, t_4, \zeta, USL, Q, \mu_0, \) and \( \delta_0) \), cost parameters \( (a_1, a_2, a_3, a_4, \) and \( C_0) \), and design parameters \( (\lambda, n, h, UCL, \) and \( H) \), the \( TC(\delta) \) can be calculated as follows:

1. Estimate the expected length of an operational cycle.
2. Calculate \( t_4 \) using Eq. (10), in which \( ATS_1 \) for any given value of shift of size \( \delta \) in the carbon emission process is calculated by a Markov chain approach [45].
3. Estimate the expected cost of an operational cycle.
4. For a given value of \( \delta \), calculate \( C_1 \) using Eq. (13).
5. Calculate \( C_2 \) using Eq. (14).
6. Calculate \( C_3 \) using Eq. (15), in which \( ATS_0 \) is calculated (\( \delta = 0 \)) by a Markov chain approach [45].
7. Calculate \( TC(\delta) \) using Eq. (17).

3.3.3. Optimization process

Fig. 1 illustrates the process of optimization design of the proposed economic-statistical \( \bar{X} \& EWMA \) scheme.

The optimization process is terminated if no further improvement in the ETC value is found. At the end of optimization process, the combination of the optimal design parameters \( (\lambda, n, h, UCL, \) and \( H) \) that ensures the minimum ETC and satisfies the constraints \( (\sigma \leq 0), (n \leq n_{\text{max}}), \) and \( (ATS_0 \geq \tau) \), is identified. Because the design is optimized under the standard condition \( (\mu_0 = 0, \sigma_0 = 1) \), the actual control limits are calculated using the actual values of \( \mu_0 \) and \( \sigma_0 \).

\[ UCL_{\text{actual}} = \mu_0 + \sigma_0 \cdot UCL \]

\[ H_{\text{actual}} = \mu_0 + \sigma_0 \cdot H \]  

(18)

A computer program using C language was developed to automate the design process of the optimal economic-statistical \( \bar{X} \& EWMA \) scheme. The program is available upon request.

4. Numerical studies

4.1. Comparison study

The effectiveness of four SPM schemes is compared in this section:

1. The basic economic-statistical \( \bar{X} \) scheme: This is a conventional economic-statistical \( \bar{X} \) scheme that uses a sample size of five \( (n = 5) \). An \( \bar{X} \) scheme is usually designed by considering a constant sample size of five [25].
2. The basic economic-statistical EWMA scheme: This EWMA scheme is designed by assuming a constant weighting parameter, \( \lambda, \) of 0.1 and a constant sample size, \( n, \) of 1. The value of parameter \( \lambda \) is subjectively selected from the widely used values of 0.05, 0.1, or 0.20 [25], and an EWMA scheme with \( n = 1 \) is known to be successful from an overall perspective [31].
3. The basic economic-statistical \( \bar{X} \& EWMA \) scheme: Similar to the basic economic-statistical EWMA scheme, this \( \bar{X} \& EWMA \) combination uses \( \lambda \) value of 0.1 and \( n \) value of 1. Following [46], the value of \( UCL \) of the \( \bar{X} \) scheme is set at 4.25, while the parameter \( H \) of the EWMA scheme is decided to ensure that the constraint of \( (ATS_0 \geq \tau) \) is satisfied.
4. The optimal economic-statistical \( \bar{X} \& EWMA \) scheme: The values of the design parameters \( (n, h, \lambda, UCL, \) and \( H) \) of this scheme are optimized by following the algorithm illustrated in Fig. 1.
To facilitate the comparison, a normalized $ETC_{normal}$ value for each SPM scheme is calculated as follows:

$$ETC_{normal} = \frac{ETC}{ETC_{opt}}$$  \hspace{1cm} (19)

where $ETC$ and $ETC_{opt}$ are the expected total costs of a specific SPM scheme and the optimal XEWMA scheme, respectively. An $ETC_{normal}$ value greater than 1 of a scheme indicates that its effectiveness is poorer than that of the optimal economic-statistical XEWMA scheme, and vice versa. The four SPM schemes are designed under the standard condition ($\mu_0 = 0$, $\sigma_0 = 1$), and $n_{max}$ is assumed to be fixed at 15 in this study, as handling a large sample size is not preferable in practice.

Because of the large number of input variables (eight process parameters $[\lambda_0, O, \mu_0, g, t_4, \zeta, USL, Q, a_1, a_2, a_3, a_4,$ and $C_E]$), the effectiveness of the four SPM schemes is investigated using a $2^{13-8}$ fractional factorial design [25]. The 13 input variables are considered as the factors, and $ETC_{normal}$ (Eq. (19)) is considered the response. Each of the 13 factors vary at two levels, as displayed in Table 2.

For each of the 32 runs resulting from the $2^{13-8}$ factorial design, the four SPM schemes are designed in such a way that each of them ensures the satisfaction of all constraints. The resultant $ETC_{normal}$ values (see Table 3) showed that the developed optimal economic-statistical XEWMA scheme consistently outperformed the other schemes throughout the 32 runs.

The average of the $ETC_{normal}$ values, $ETC_{normal}$, over the 32 runs for each scheme, was also calculated. The values of $ETC_{normal}$ showed that from a general viewpoint (over different combinations of $\lambda_0$, $O$, $\mu_0$, $g$, $t_4$, $\zeta$, $USL$, $Q$, $a_1$, $a_2$, $a_3$, $a_4$, and $C_E$), the optimal economic-statistical XEWMA scheme outperformed (in terms of $ETC$) the basic economic-statistical, basic economic-statistical EWMA, and basic economic-statistical XEWMA schemes by about 40%, 77%, and 28%, respectively. The improvement in the effectiveness of the optimal economic-statistical XEWMA scheme compared with that of the other three schemes was further investigated using paired $t$-tests [25] (see bottom row of Table 3). The results showed that the improvements in the effectiveness of the optimal economic-statistical XEWMA scheme compared with the basic economic-statistical $X$ scheme (p-value = 0.004), basic economic-statistical EWMA scheme (p-value = 0.010), and basic economic-statistical XEWMA scheme (p-value = 0.012) were all statistically significant, using a significance level of 5%.

### 4.2 Sensitivity analysis

The impacts of the 13 input variables ($\lambda_0$, $O$, $\mu_0$, $g$, $t_4$, $\zeta$, $USL$, $Q$, $a_1$, $a_2$, $a_3$, $a_4$, and $C_E$) on the response parameter ($ETC$) of the optimal

---

**Table 2**

<table>
<thead>
<tr>
<th>Input factor</th>
<th>Low level</th>
<th>High level</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\lambda_0$: Rate of occurrence of the assignable cause (occurrences per month)</td>
<td>0.5</td>
<td>1.0</td>
</tr>
<tr>
<td>$Q$: Maximum allowable inspection rate (number of data inspected per month)</td>
<td>2</td>
<td>5</td>
</tr>
<tr>
<td>$O$: Mean of the mean shifts $\delta$ in the amount of carbon emission process</td>
<td>0.5</td>
<td>3.5</td>
</tr>
<tr>
<td>$g$: Time to estimate and test an observed data of a sample of carbon emission (month)</td>
<td>0.001388</td>
<td>0.006944</td>
</tr>
<tr>
<td>$t_4$: Time period from the detection of the lack of control to the location and removable of the assignable cause (month)</td>
<td>0.034</td>
<td>0.10</td>
</tr>
<tr>
<td>$\zeta$: Minimum allowable in-control ATS$_0$ (month)</td>
<td>300</td>
<td>800</td>
</tr>
<tr>
<td>$USL$: Upper specification limit (i.e. carbon-cap) of the amount of carbon emission (tons per month)</td>
<td>$3\sigma_0$</td>
<td>$6\sigma_0$</td>
</tr>
<tr>
<td>$Q$: Amount of carbon emission (tons per month)</td>
<td>5,000,000</td>
<td>70,000,000</td>
</tr>
<tr>
<td>$a_1$: Fixed component of sampling cost ($)</td>
<td>200</td>
<td>500</td>
</tr>
<tr>
<td>$a_2$: Variable component of sampling cost ($)</td>
<td>20</td>
<td>50</td>
</tr>
<tr>
<td>$a_3$: Cost of finding and fixing an assignable cause ($)</td>
<td>500</td>
<td>1000</td>
</tr>
<tr>
<td>$a_4$: Cost of examining a false alarm ($)</td>
<td>1000</td>
<td>2000</td>
</tr>
<tr>
<td>$C_E$: Average penalty cost for an out-of-specification amount of carbon emission ($ per ton)</td>
<td>100</td>
<td>200</td>
</tr>
<tr>
<td>Run</td>
<td>Values of the input factors</td>
<td>( ET_{\text{normal}} )</td>
</tr>
<tr>
<td>---</td>
<td>---</td>
<td>---</td>
</tr>
<tr>
<td>1</td>
<td>1.0</td>
<td>2</td>
</tr>
<tr>
<td>2</td>
<td>1.0</td>
<td>5</td>
</tr>
<tr>
<td>3</td>
<td>0.1</td>
<td>2</td>
</tr>
<tr>
<td>4</td>
<td>0.5</td>
<td>2</td>
</tr>
<tr>
<td>5</td>
<td>0.5</td>
<td>5</td>
</tr>
<tr>
<td>6</td>
<td>1.0</td>
<td>5</td>
</tr>
<tr>
<td>7</td>
<td>1.0</td>
<td>5</td>
</tr>
<tr>
<td>8</td>
<td>1.0</td>
<td>2</td>
</tr>
<tr>
<td>9</td>
<td>0.5</td>
<td>2</td>
</tr>
</tbody>
</table>

**Table 3**

Comparison of the four schemes in the 2\(^3\) \(x^2\) experiment.

ETC = ETC of a chart over 32 runs - ETC of the optimal X&EWMA scheme over 32 runs. Positive values indicate superiority of optimal X&EWMA scheme to other schemes.
economic-statistical X&EWMA scheme were also investigated using the $2^{13}\times 8$ factorial design indicated in Table 2. Because the replication size was 1, the higher order (higher than or equal to the third order) interaction effects were combined to estimate the sum of squares of the error. The significant main and two-factor interaction effects were identified by an analysis of variance (ANOVA). Before performing the ANOVA test, a normality test of the ETC data was performed to check the model adequacy. The data on ETC were initially not normal; therefore, Johnson transformation was conducted before performing the ANOVA test (see Fig. 2).

The results of the ANOVA test, as shown in Table 4, confirm that only four main factor effects (bold text) were statistically significant. As shown in Table 4, the ETC of the optimal economic-statistical X&EWMA scheme is positively affected by $\mu_d$ ($p$-value = 0.028), $Q$ ($p$-value = 0.001), and $C_K$ ($p$-value = 0.010). This implies that a larger $\mu_d$ or $Q$ or $C_K$ value can result in a larger ETC, and vice versa. Conversely, the ETC is negatively affected by $USL$ ($p$-value = 0.002). This means a tighter $USL$ (carbon-cap) can result in a larger ETC, and vice versa. This is justifiable as a smaller $USL$ needs to utilize more resources and more investigations.

5. Case study

The design and application of the optimal economic-statistical X&EWMA scheme are demonstrated based on real data on the amount of carbon emissions from factories in the United States and are explained in the following steps.

5.1. Data collection

In 2017, the estimated GHGs emissions from the industrial sector represented 22.2% of the total emissions of GHGs in the United States [47]. Manufacturing and industrial processes together produce large amounts of GHGs specifically $CO_2$. The State Department of Environmental Conservation (DEC) of New York, as part of its mission to conserve natural resources and protect the environment, keeps records of different sources of pollution, including industrial facilities that emit or have the potential to emit air pollutants, requiring these facilities to

Table 4

<table>
<thead>
<tr>
<th>Input factors</th>
<th>Effects on the ETC of the optimal X&amp;EWMA scheme</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\lambda$</td>
<td>0.2006</td>
</tr>
<tr>
<td>$Q$</td>
<td>1.5091</td>
</tr>
<tr>
<td>$a_1$</td>
<td>-0.0577</td>
</tr>
<tr>
<td>$a_2$</td>
<td>-0.0410</td>
</tr>
<tr>
<td>$a_3$</td>
<td>0.0704</td>
</tr>
<tr>
<td>$a_4$</td>
<td>0.0933</td>
</tr>
<tr>
<td>$C_K$</td>
<td>0.5796</td>
</tr>
<tr>
<td>$\mu_d$</td>
<td>0.3916</td>
</tr>
<tr>
<td>$g$</td>
<td>-0.1746</td>
</tr>
<tr>
<td>$t_4$</td>
<td>0.1384</td>
</tr>
<tr>
<td>$\zeta$</td>
<td>-0.1001</td>
</tr>
<tr>
<td>$USL$</td>
<td>-0.9402</td>
</tr>
<tr>
<td>$a_{11}$</td>
<td>0.0577</td>
</tr>
<tr>
<td>$a_{12}$</td>
<td>0.0129</td>
</tr>
<tr>
<td>$a_{13}$</td>
<td>-0.0208</td>
</tr>
<tr>
<td>$a_{14}$</td>
<td>0.0189</td>
</tr>
<tr>
<td>$a_{21}$</td>
<td>0.0516</td>
</tr>
<tr>
<td>$a_{22}$</td>
<td>0.0129</td>
</tr>
<tr>
<td>$a_{23}$</td>
<td>-0.0208</td>
</tr>
<tr>
<td>$a_{24}$</td>
<td>0.0189</td>
</tr>
<tr>
<td>$a_{31}$</td>
<td>0.0516</td>
</tr>
<tr>
<td>$a_{32}$</td>
<td>0.0129</td>
</tr>
<tr>
<td>$a_{33}$</td>
<td>-0.0208</td>
</tr>
<tr>
<td>$a_{34}$</td>
<td>0.0189</td>
</tr>
<tr>
<td>$a_{41}$</td>
<td>0.0516</td>
</tr>
<tr>
<td>$a_{42}$</td>
<td>0.0129</td>
</tr>
<tr>
<td>$a_{43}$</td>
<td>-0.0208</td>
</tr>
<tr>
<td>$a_{44}$</td>
<td>0.0189</td>
</tr>
</tbody>
</table>

Fig. 2. Normality check of the ETC data of the optimal economic-statistical X&EWMA scheme.
report their emissions of pollutants. These reports are public information and can be obtained from the permitting authority [48]. The DEC monitors these industrial facilities to ensure that the source complies with the emission limit, or other pollution control requirements. The SPM charts would be appropriate to achieve this objective.

A dataset of the annual CO$_2$ emissions measured in tons for 306 facilities at 53 different counties in the New York State in 2011 [49], collected by the State DEC, has been utilized in this study. Based on the annual data obtained, the monthly CO$_2$ emissions data have been calculated and used for illustrating the concept of SPM schemes for monitoring carbon emissions and controlling air quality.

5.2. Model adequacy test

In designing an SPM scheme for variable-type quality characteristics, the quality characteristic $x$ (data on the amount of carbon emissions in this study) is presumed to be normally and independently distributed. A slight or moderate degree of violation of the normality assumption may not affect the effectiveness of an SPM scheme. However, a slight dependency (autocorrelation) among the data significantly affects the effectiveness of an SPM scheme, and thus the dependency of the data should be checked before designing. To verify the assumption of independency, a time series plot was used to represent the carbon emissions data as in Fig. 3(a), which does not show any evidence of seasonality in the data. In addition, autocorrelation function (ACF) and partial ACF (PACF) were also drawn to explore how the data points are related to each other (see Fig. 3(b-c)). Fig. 3(a-c) confirms that the emission data are independent. However, according to the normal probability plot shown in Fig. 4(a), the data are not normally distributed ($p$-value < 0.01). Thus, a transformation technique is required to transform the non-normally distributed data into normally distributed data [50]. The ordered quantile (ORQ) normalization technique was used for this purpose, achieved using the package “bestNormalize” (version 1.4.2) available in R programming language (version 3.6.2). The transformed data satisfied the normality assumption ($p$-value > 0.15), as illustrated in Fig. 4(b). Finally, the concept of SPM schemes for monitoring carbon emissions was demonstrated based on the transformed data that satisfied both the normality and independency assumptions.

5.3. Design and application of the proposed SPM scheme

The design of an SPM scheme is accomplished in two phases: Phase I and Phase II operations. In Phase I operation, at least 25–30 samples,
each of size five, are usually recommended in designing a classical $X$ scheme [25]. The objective of collecting samples in Phase I is to estimate the IC values of $\mu_0$ and $\sigma_0$ for designing the SPM scheme. In Phase II, the SPM scheme designed at the end of Phase I is utilized for monitoring the process in the future.

5.3.1. Phase I operation

After the initial screening, 295 observations (59 samples with sample size of five) of carbon emissions data were used in Phase I for designing the basic $X$ scheme. Fig. 5 shows phase I SPM scheme. Designing both the $X$ scheme (for monitoring the process mean) and $R$ scheme (for monitoring the process dispersion) is recommended in Phase I to ensure that no assignable cause is presented in the process (i.e., the process is in the IC state), and that the estimated $\mu_0$ and $\sigma_0$ that will be used in Phase II are consistent [25].

Fig. 5 shows that all the 59 sample points are plotted within the control limits ($X$ scheme: $UCL = 1.357$, $CL = 0.015$, $LCL = -1.327$; $R$ scheme: $UCL = 4.918$, $CL = 2.326$, $LCL = 0$) of the $X$&$R$ scheme, indicating that the carbon emission process is in IC state ($\mu_0 = 0.015$ and $\sigma_0 = 0.9995$).

5.3.2. Phase II operation

The process parameters in the IC state ($\mu_0 = 0.015$, $\sigma_0 = 0.9995$),

Fig. 4. Normality check of the carbon emission data.

Fig. 5. $X$&$R$ scheme for carbon emission data in Phase I.
estimated in Phase I, were used in this phase to design a basic economic-statistical $X$ scheme, a basic economic-statistical EWMA scheme, a basic economic-statistical $X$&EWMA scheme, and an optimal economic-statistical $X$&EWMA scheme for monitoring the forthcoming data on the amount of carbon emissions. To design and demonstrate the effectiveness of the proposed SPM schemes, the other required design parameter values assumed are as follows:

- $\lambda_0$ (occurrence rate of the assignable cause, occurrences per month) = 0.01
- $(\text{maximum allowable inspection rate (per month)}) = 5$
- $\mu_1$ (mean of the $d$ values in the emission process) = 0.75
- $g$ (time to estimate and test the observed data of a sample of carbon emissions, month) = 0.00035
- $t_d$ (time period from the detection of an OOC state to the identification and fixation of the assignable cause, month) = 0.1
- $\zeta$ (minimum allowable IC ATSD, month) = 400
- USL (upper specification limit (carbon-cap) of the amount of carbon emissions, tons per month) = $4t_0$
- $Q$ (average amount of carbon emissions from an industry, tons per month) = 12500
- $a_1$ (fixed part of the sampling cost, $) = 0.5$
- $a_2$ (variable part of the sampling cost, $) = 0.1$
- $a_3$ (cost of identifying and fixing an assignable cause, $) = 100$
- $a_4$ (cost of investigating a false alarm, $) = 200$
- $C_p$ (average penalty cost for an out-of-specification [out-of-carbon cap] amount of carbon emissions, $ per ton of CO$_2$ emissions) = 150

The abovementioned hypothetical data have been used in this study to illustrate the effectiveness of the proposed SPM schemes. However, the value of the penalty cost $C_p$ was set at $150/ton, according to the current carbon tax in Sweden [51]. In addition, real data for other parameters can be used when the data are publicly accessible. In this example, the maximum allowable sample size $n_{max}$ is considered 10. The developed computer program used to design the four SPM schemes and the parameter values of each scheme are listed below. Note that the designs of all SPM schemes ensure that all constraints (4) and (5) in Section 3 are satisfied.

- Basic economic-statistical $X$ scheme:

  \[ n = 5, h = 1.0, UCL = 1.2697, ETC = 20817.98, ETC_{\text{normal}} = 2.607 \]

- Basic economic-statistical EWMA scheme:

  \[ n = 1, h = 0.20, \lambda = 0.10, H = 0.7539, ETC = 12862.38, ETC_{\text{normal}} = 1.610 \]

- Basic economic-statistical $X$&EWMA scheme:

  \[ n = 1, h = 2.0, \lambda = 0.10, UCL = 4.2629, H = 0.7552, ETC = 12919.80, ETC_{\text{normal}} = 1.618 \]

- Optimal economic-statistical $X$&EWMA scheme:

  \[ n = 10, h = 2.0, \lambda = 0.17, UCL = 1.1725, H = 0.2559, ETC = 7986.75, ETC_{\text{normal}} = 1.000 \]

Because of the unavailability of the real data on OOC states, the effectiveness of the four designed SPM schemes for detecting an OOC signal was investigated via simulation study, in which 20 sample data on the amount of carbon emissions were generated. The first 10 sample data were simulated under the IC condition, and the other 10 considering a 1.0 shift in the mean of the carbon emission process (i.e., under the OOC condition). It is a common practice in the literature to use simulation to study the effectiveness of a proposed model when real data are unavailable (for instance, see [28,29,52]). All the 20 simulated data are plotted on the four SPM schemes, as shown in Fig. 6.

As shown in Fig. 6(a-d), all three basic SPM schemes ($X$, EWMA, and $X$&EWMA schemes) were unable to identify the OOC condition of the process. However, the proposed optimal $X$&EWMA scheme identified the OOC condition by the 13th sample, evidently demonstrating its supremacy over the basic SPM counterparts. The improvement in the detection effectiveness results in overall cost savings (in terms of ETC) was about 160%, 61%, and 62%, compared to the basic X, basic EWMA, and basic $X$&EWMA schemes, respectively, in this study.

6. Conclusions

The reduction of GHGs emissions is considered as a major issue within the global community. Amongst all the GHGs, CO$_2$ is considered as the most significant contributor to the changes in global climactic conditions. Hence, researchers and professionals have intensely focused on finding suitable methods for monitoring and controlling CO$_2$ emissions. The industrial sector is one of the fastest-growing sources of GHGs, due to the excessive consumption of energy required to cope with the growing production of energy exhaustive products. The continuous monitoring of CO$_2$ emissions from different industrial facilities can be an important step in reducing carbon emissions and encouraging them to use cleaner energy. This article presents an optimal economic-statistical design of the combined $X$&EWMA scheme for efficient monitoring of the carbon emissions from industrial facilities. The design of the proposed SPM scheme is based on emissions data collected from different industrial facilities. However, the data can also be collected from only one location, if the focus is to monitor and control a single facility. The effectiveness of the proposed optimal SPM scheme was compared with that of other monitoring schemes, namely the basic $X$, basic EWMA, and basic $X$&EWMA schemes. The comparison study showed that the proposed optimal $X$&EWMA scheme reduced the expected total cost incurred owing to carbon emissions and operation of the SPM scheme by about 40%, 79%, and 29%, compared with the basic $X$, basic EWMA, and basic $X$&EWMA schemes, respectively. Finally, the design and application of the proposed SPM scheme are illustrated based on real data carbon emissions collected from different industrial facilities. The same SPM scheme can also be used for monitoring the emissions from other facilities in other sectors, such as transportation, building and construction, and agriculture.

In this study, the random shift in the carbon emission process is modeled by a Rayleigh distribution. In future study, the effectiveness of the proposed SPM scheme can be investigated over other distributions of the shift, such as uniform or beta distribution. Other SPM scheme such as dual-EWMA or $X$&CUSUM scheme can also be designed for monitoring the emission process, and the performance of these schemes can be compared with that of the optimal $X$&EWMA scheme proposed in this study.

CRediT authorship contribution statement

Mohammad Shamsuzzaman: Conceptualization, Methodology, Software, Writing – original draft, Funding acquisition. Ahm Sham: Visualization, Validation, Writing – review & editing. Ahmed Maged: Visualization, Writing – review & editing. Salah Haridy: Methodology, Software, Writing – original draft. Hamdi Bashir: Investigation, Validation, Writing – review & editing. Azharul Karim: Validation, Writing – review & editing.

Declaration of Competing Interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence
the work reported in this paper.

Acknowledgements

This article is a part of an ongoing research project titled "statistical monitoring of carbon emissions in industry using the Shewhart-EWMA scheme," conducted by the Sustainable Engineering Asset Management research group at the University of Sharjah, UAE.

Funding

This work is supported by the University of Sharjah Competitive Research Grant (Project No. 200204051161).

References


Duncan AJ. The economic design of X charts used to maintain current control of a process. J Am Stat Assoc 1956;51(274):228–42.

Chen YS, Yang YM. Economic design of control charts with weibull in-control times when there are multiple assignable causes. Int J Prod Econ 2002;77(1):17–23.


